Design and Analysis of Experiments 04 - Interpretation of results

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Outline

Interpretation of results

- Comparison among treatment means
- Graphical comparison of means
- Contrasts
- Comparing pairs of treatment means
- Comparing treatment means with control

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Interpretation of results

- After conducting the experiment, performing the statistical analysis, and investigating the underlying assumptions, the experimenter is ready to draw practical conclusions about the problem under study.
- Often this is relatively easy, particularly in simple experiments, where conclusions might be drawn informally, perhaps by inspecting graphical displays such as box plots and scatter diagrams.
- However, in some cases, more formal techniques need to be applied.

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Regression Model

- Factors in an experiment can be either quantitative or qualitative.
- Quantitative factors: levels associated with points on a numerical scale (e.g., temperature, time).
- Qualitative factors: levels without meaningful numerical order (e.g., operators, batches).
- Both types are initially treated identically in design and analysis.

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Prediction and Interpolation

- When dealing with quantitative factors, the experimenter is often interested in predicting responses at intermediate levels.
- For instance, with levels of 1.0, 2.0, and 3.0 hours, predicting the response at 2.5 hours may be desired.
- This involves developing an empirical model, often using regression analysis.

Regression Analysis

- Regression analysis is used to fit empirical models to the data.
- The method of least squares is commonly applied to estimate parameters.

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Scatter Diagrams and Model Fitting

Figure 3.10 presents scatter diagrams of etch rate y versus power x for the experiment in Example 3.1. A strong relationship is observed. We can first attempt to fit a linear model:

$$y = \beta_0 + \beta_1 x + \epsilon$$

where β_0 and β_1 are unknown parameters and ϵ is a random error term. The least squares fit yields:

$$y = 137.62 + 2.527x$$

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Quadratic Model

At higher power settings, the linear model may not be satisfactory. A quadratic term can be added:

$$y = 1147.77 - 8.2555x + 0.028375x^2$$

This quadratic model provides a better fit for the data at higher power levels.

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Scatter Diagrams and Model Fitting



FIGURE 3.10 Scatter diagrams and regression models for the etch rate data of Example 3.1

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Comparison among treatment means

- Analysis of variance for fixed effects model
- Rejection of null hypothesis implies differences between treatment means
- Further comparisons and analysis may be useful

Definitions

- *i*th treatment mean: μ_i
- Estimate of μ_i : \bar{y}_i
- Comparisons made using treatment totals $\{y_{i.}\}$ or averages $\{\bar{y}_{i}\}$

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Multiple Comparison Methods

- Used for comparing individual treatment means or groups of means
- Various methods available (to be discussed in following sections)

Graphical Comparisons of Means

- Easy to develop graphical procedure
- Consider factor with *a* levels and treatment averages $\{\bar{y}_i\}$
- If all means are identical, $\{\bar{y}_i\}$ should behave like random samples from $N(\mu, \sigma/\sqrt{n})$
- Replace unknown σ with $\sqrt{MS_E}$ from ANOVA
- Use t-distribution with scale factor $\sqrt{MS_E/n}$
- Plot treatment means and overlay scaled t-distribution

Graphical Comparisons of Means



FIGURE 3.11 Etch rate averages from Example 3.1 in relation to a *t* distribution with scale factor $\sqrt{MS_E/n} = \sqrt{330.70/5} = 8.13$

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Example: Etch Rate Data

- Refer to Figure 3.11 (not included in this presentation)
- $\bullet\,$ t-distribution origin at 615 Å/min
- Scale factor: $\sqrt{MS_E/n} = \sqrt{330.70/5} = 8.13$

Interpretation of Graphical Method

- No single position where all means fit the distribution
- Implies not all means are equal (confirms ANOVA results)
- Suggests all four power levels (160, 180, 200, 220 W) produce different mean etch rates

Conclusion

- Graphical method is simple but effective
- Useful for many multiple comparison problems
- More formal methods are available (to be discussed next)

Introduction to Contrasts

- Many multiple comparison methods use the idea of a contrast
- Example: Plasma etching experiment from Example 3.1
- Goal: Identify which power settings cause differences in etch rates

Hypothesis Testing with Contrasts

• Hypothesis 1: 200 W and 220 W produce the same etch rate

 $H_0: \mu_3 = \mu_4$ $H_1: \mu_3 \neq \mu_4$

• Hypothesis 2: Average of lowest levels = Average of highest levels

$$H_0: \mu_1 + \mu_2 = \mu_3 + \mu_4$$
$$H_1: \mu_1 + \mu_2 \neq \mu_3 + \mu_4$$

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Contrasts

Definition of a Contrast

• A contrast is a linear combination of parameters:

$$\sum_{i=1}^{a} c_i \mu_i$$

• Contrast constants c_1, c_2, \ldots, c_a sum to zero: $\sum_{i=1}^a c_i = 0$

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Expressing Hypotheses as Contrasts

$$H_0: \sum_{i=1}^{a} c_i \mu_i = 0$$
 vs $H_1: \sum_{i=1}^{a} c_i \mu_i \neq 0$

• For $H_0: \mu_3 = \mu_4$

$$c_1 = c_2 = 0, c_3 = +1, c_4 = -1$$

• For $H_0: \mu_1 + \mu_2 = \mu_3 + \mu_4$

$$c_1 = c_2 = +1, c_3 = c_4 = -1$$

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Testing Contrasts: Method 1 (t-test)

• Express contrast in terms of treatment averages:

$$C=\sum_{i=1}^{a}c_{i}\bar{y}_{i.}$$

• Variance of C (for equal sample sizes):

$$V(C) = \sigma^2 \frac{\sum_{i=1}^{a} c_i^2}{n}$$

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Testing Contrasts: t-statistic

• If H_0 is true, the ratio follows a t-distribution:

$$t_0 = \frac{C}{\sqrt{MS_E \frac{\sum_{i=1}^a c_i^2}{n}}}$$

- MS_E is the mean square error from ANOVA
- Degrees of freedom: N a

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Conclusion

- Contrasts are powerful tools for multiple comparisons
- They allow testing specific hypotheses about differences among treatments
- t-test method is one way to test contrasts
- Other methods exist (to be discussed in future lectures)

Testing Contrasts: t-test Method (continued)

- Replace unknown variance σ^2 with MSE
- Use the test statistic:

$$t_0 = \frac{\sum_{i=1}^{a} c_i \bar{y}_{i.}}{\sqrt{MSE \frac{\sum_{i=1}^{a} c_i^2}{n}}}$$

• Reject
$$H_0$$
 if $|t_0| > t_{lpha/2, N-a}$

Testing Contrasts: F-test Method

- Square of a t random variable with ν df is an F random variable with 1 numerator df and ν denominator df
- F-statistic:

$$F_{0} = t_{0}^{2} = \frac{(\sum_{i=1}^{a} c_{i} \bar{y}_{i.})^{2}}{MSE \frac{\sum_{i=1}^{a} c_{i}^{2}}{n}}$$

• Reject H_0 if $F_0 > F_{\alpha,1,N-a}$

Contrast Sum of Squares

• F-statistic can be written as:

$$F_0 = \frac{MSC}{MSE}$$

• Where the contrast sum of squares is:

$$SSC = \frac{(\sum_{i=1}^{a} c_i \bar{y}_{i.})^2}{\frac{1}{n} \sum_{i=1}^{a} c_i^2}$$

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Graphical Comparisons of Means

TABLE 3.11

Analysis of Variance for the Plasma Etching Experiment

Source of Variation	Sum of Squares	Degrees of Freedom	Mean Square	F_{0}	P-Value
Power setting Orthogonal contrasts	66,870.55	3	22,290.18	66.80	< 0.001
$C_1: \mu_1 = \mu_2$	(3276.10)	1	3276.10	9.82	< 0.01
$C_2: \mu_1 + \mu_3 = \mu_3 + \mu_4$	(46,948.05)	1	46,948.05	140.69	< 0.001
$C_3: \mu_3 = \mu_4$	(16,646.40)	1	16,646.40	49.88	< 0.001
Error	5,339.20	16	333.70		
Total	72,209.75	19			

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Confidence Interval for a Contrast

- Consider the contrast: $\sum_{i=1}^{a} c_i \mu_i$
- Estimate using sample means: $\sum_{i=1}^{a} c_i \bar{y}_{i.}$
- Variance of the estimate:

$$V(C) = \frac{\sigma^2}{n} \sum_{i=1}^{a} c_i^2$$

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Confidence Interval Formula

• $100(1-\alpha)\%$ confidence interval:

$$\sum_{i=1}^{a} c_i \bar{y}_{i.} - t_{\alpha/2, N-a} \sqrt{\frac{MSE}{n} \sum_{i=1}^{a} c_i^2} \le \sum_{i=1}^{a} c_i \mu_i$$
$$\le \sum_{i=1}^{a} c_i \bar{y}_{i.} + t_{\alpha/2, N-a} \sqrt{\frac{MSE}{n} \sum_{i=1}^{a} c_i^2}$$

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Summary

- Two methods for testing contrasts: t-test and F-test
- Both methods use the mean square error (MSE) from ANOVA
- Confidence intervals provide a range of plausible values for the contrast
- These methods allow for specific comparisons among treatment means

Comparing pairs of treatment means

- Often need to compare pairs of means after ANOVA
- Interest in contrasts of the form $\mu_i \mu_j$ for all $i \neq j$
- Multiple methods available; we'll focus on Tukey's test

Pairwise Comparisons

- Null hypotheses: $H_0: \mu_i = \mu_j$ for all $i \neq j$
- Alternative hypotheses: $H_1: \mu_i \neq \mu_j$
- \bullet Goal: Control overall significance level at α

Tukey's Test

- Proposed by Tukey (1953)
- Controls experimentwise (family) error rate at α
- Exact when sample sizes are equal
- Conservative (at most α) when sample sizes are unequal

Studentized Range Statistic

• Tukey's test uses the studentized range statistic:

$$q = \frac{y_{\max} - y_{\min}}{\sqrt{MSE/n}}$$

- y_{max} and y_{min} are the largest and smallest sample means
- Distribution values found in Appendix Table VII

Tukey's Test: Equal Sample Sizes

• Declare means significantly different if:

$$|y_{i.} - y_{j.}| > T_{\alpha} = q_{\alpha}(a, f) \sqrt{\frac{MSE}{n}}$$

• Confidence intervals:

$$y_{i.} - y_{j.} - T_{\alpha} \leq \mu_i - \mu_j \leq y_{i.} - y_{j.} + T_{\alpha}$$

- $q_{\alpha}(a, f)$ is the upper α percentage point of q
- a is the number of means, f is degrees of freedom for MSE

Tukey's Test: Unequal Sample Sizes

- Known as Tukey-Kramer procedure
- Declare means significantly different if:

$$|y_{i.}-y_{j.}| > q_{\alpha}(a,f) \sqrt{\frac{MSE}{2}\left(\frac{1}{n_i}+\frac{1}{n_j}\right)}$$

• Confidence intervals:

$$y_{i.} - y_{j.} - T_{lpha} \le \mu_i - \mu_j \le y_{i.} - y_{j.} + T_{lpha}$$

where $T_{lpha} = q_{lpha}(a, f) \sqrt{\frac{MSE}{2} \left(\frac{1}{n_i} + \frac{1}{n_j}\right)}$

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Advantages of Tukey's Test

- Controls experimentwise error rate
- Provides confidence intervals
- Suitable for all pairwise comparisons
- Relatively easy to compute and interpret
- Widely available in statistical software

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Summary

- Tukey's test is useful for all pairwise comparisons
- Controls overall error rate at α
- Uses studentized range statistic
- Can be applied to equal or unequal sample sizes
- Provides both hypothesis tests and confidence intervals

Comparing treatment means with control

- In many experiments, one treatment is a control
- Goal: Compare each of the other *a* − 1 treatment means with the control
- Only a 1 comparisons are needed
- Dunnett (1964) developed a procedure for this scenario

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Dunnett's Procedure: Setup

- Assume treatment *a* is the control
- We want to test the hypotheses:

 $H_0: \mu_i = \mu_a$ $H_1: \mu_i \neq \mu_a$

for i = 1, 2, ..., a - 1

Dunnett's Procedure: Test Statistic

• Compute the observed differences in sample means:

$$|\bar{y}_{i.} - \bar{y}_{a.}|$$
 for $i = 1, 2, \dots, a - 1$

• Reject $H_0: \mu_i = \mu_a$ if:

$$|ar{y}_{i.} - ar{y}_{a.}| > d_{lpha}(a-1,f)\sqrt{MSE\left(rac{1}{n_i} + rac{1}{n_a}
ight)}$$

• $d_{\alpha}(a-1, f)$ is given in Appendix Table VIII

• f is the degrees of freedom associated with MSE

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Dunnett's Procedure: Key Points

- Modification of the usual t-test
- α is the joint significance level for all a-1 tests
- Controls the family-wise error rate
- Both two-sided and one-sided tests are possible
- More powerful than Tukey's test for this specific scenario

Confidence Intervals

- Dunnett's procedure can also be used to construct confidence intervals
- $100(1-\alpha)\%$ confidence intervals:

$$(\bar{y}_{i.}-\bar{y}_{a.})\pm d_{lpha}(a-1,f)\sqrt{MSE\left(rac{1}{n_{i}}+rac{1}{n_{a}}
ight)}$$

• Simultaneous confidence level is $100(1-\alpha)\%$ for all a-1 intervals

Advantages of Dunnett's Procedure

- Specifically designed for comparisons with a control
- More powerful than general pairwise comparison methods
- Controls family-wise error rate
- Provides both hypothesis tests and confidence intervals
- Widely available in statistical software

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When to Use Dunnett's Procedure

- Experiments with a clear control group
- When all comparisons are against the control
- No interest in comparisons between non-control treatments
- Examples:
 - Drug trials comparing new treatments to a placebo
 - Agricultural experiments comparing new varieties to a standard

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Summary

- Dunnett's procedure is optimal for comparing treatments with a control
- Modifies t-test to account for multiple comparisons
- Controls family-wise error rate at α
- Provides both hypothesis tests and confidence intervals
- More powerful than general pairwise comparison methods for this specific scenario